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Submission date: 15-Apr-2023 01:36PM (UTC+0700) Submission ID: 2065133818 File name: e\_Openness,\_and\_Economic\_Growth\_in\_Southeast\_Asian\_Countries.pdf (412.75K) Word count: 4034 Character count: 22150 Pertanika J. Soc. Sci. & Hum. 25 (S): 175 - 186 (2017)



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# Environmental Degradation, Trade Openness, and Economic Growth in Southeast Asian Countries

Shanty Oktavilia<sup>1,2\*</sup>, Franciscus Xaverius Sugiyanto<sup>1</sup>, Firmansyah Firmansyah<sup>1</sup> and Wahyu Widodo<sup>1</sup>

<sup>1</sup>Economics, Diponegoro University, FEB UNDIP, Jl Prof Sudharto <mark>532 Tembalang, Semarang, Indonesia</mark> <sup>2</sup>Economics Department, Semarang State University FE UNNES, Jl. Sekaran, Gunung Pati, Semarang, Indonesia

### ABSTRACT

The trade-off between economic grow and environmental quality has long been debated. The aim of this tudy was to find out if the hypothesis of the Environmental Kuznets Curve (EKC) in five Southeast Asian countries - Indonesia, Malaysia, the Philippines, Singapore, and Thailand – is supported. This study analysed the effect of GDP per-capita and the ratio of trade openness on  $CO_2$  emissions. Using 2 nual data from 1975 to 2014, this study employed the Error Correction Model (ECM) 19 test the EKC hypothesis for each country and applied the fixed effect panel data model to test the EKC hypothesis for all countries. The results showed that in the long run, the inverted U of the EKC hypothesis was supported in Singapore, without any turning point in Indonesia and in the Philippines, but no evidence was found for Malaysia are philailand. Except for Singapore, trade liberalisation had positive effect on  $CO_2$  emissions. The results of panel data analysis indicated that, in general, the inverted U of the EKC hypothesis was supported in Southeast Asian countries, while trade liberalisation positively affected the increase of  $CO_2$  emissions.

*Keywords:* CO<sub>2</sub> emission, economic growth, environmental degradation, error correction model, EKC hypothesis, panel data, trade openness

ARTICLE INFO

Article history: Received: 20 May 2017 Accepted: 01 October 2017

E-mail addresses: oktavilia@mail.unnes.ac.id (Shanty Oktavilia) firmansyah@live.undip.ac.id (Firmansyah Firmansyah) fxsugiyanto09@gmail.com (Franciscus Xaverius Sugiyanto) wahyuwid2002@live.undip.ac.id (Wahyu Widodo) \* Corresponding author

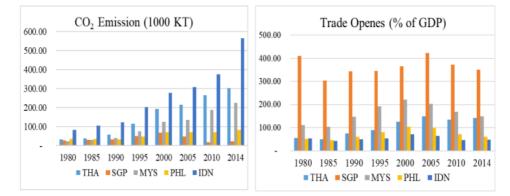
ISSN: 0128-7702 © Universiti Putra Malaysia Press

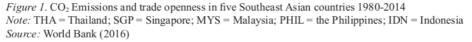
### INTRODUCTION

Economic growth is the focus of many developing countries but it also leads to negative externalities such as environmental degradation. Todaro (2000) points out that this condition may be a result of

exploitation and depletion of the natural resources. For example, economic openness of economic stimulates the economic transformation from agriculture to industry and is suspected to be the cause of the environmental degradation (Huang & Labys, 2002). Moreover, industrialisation tends to increase the CO<sub>2</sub> emission as the use of energy increases (Ekbom & Dahlberg, 2008).

Several developing countries in Southeast Asia whose national development depend on the existence of their natural resources placeless attention on the environmental sustainability issues. The link between trade openness and economic growth in Southeast Asian Countries during 1980 – 2014 is shown in Figure 1.





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The relationship between rate of economic growth and environmental degradation can be explained using Environmental Kuznets Curve hypothesis. This hypothesis explains that when a country's income remains relatively low, the attention of the government will be focused on how to increase the national income through economic growth. The long-run interaction between economic growth and environmental quality resembles an inverted U-curve (Grossman & Krueger, 1991), and this phenomenon is called Environmental Kuznets Curve -EKC.

In the developing countries, economic growth which leads an increase in population often ignores the environmental quality issues. Meanwhile, in the developed countries, the awareness of environment quality has increased as their income has improved in which the environmental degradation and the environmental pollution will be compensated. The demand of the people on environmental quality that

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could be improved by improving social supervision and government regulation aims to achieve the prosperity of the people (Mason & Swanson 2003).

Empirically, the trade-off between economic growth and environmental degradation has become agglemic issues. Researchers have analysed the relationship between environmental quality, economic growth, and liberalisation policies. Given this situation, Huang and Labys (2002) identify the argument focusing on two interrelated issues. The first is following the Kyoto Protocol Agenda of the trend in increasing carbon emission. In the future, one of the most important challenges for environmental policies is the effort to reduce CO<sub>2</sub> emission; it has been realized that formulating the relationship between CO<sub>2</sub> emission and economic growth for the benefit of the public policy is crucial. The second is the trade openness to accelerate the economic growth; yet, increase the pollution level. The international trade as a consequence of globalisation is characterised by economy growth; however, environmental pollution will also potentially increase.

This study examines the influence of growth and trade openness on the environmental quality in Southeast Asian countries by employing the error correction model (ECM), and using panel data regression. Gata from 1975 to 2014 of five South East Asian countries, including Indonesia, Malaysia, Singapore, Thailand, and the Philippines was obtained and analysed.

### LITERATURE REVIEW

# Environmental Degradation and

The Environmental Kuznets Curve (EKC) due to its resemblance to Kuznets gerve, which is an inverted U-shape curve, can be utilised to describe the relationship between inequality of income and economic growth (Huang & Labys, 2002; Panayotou, 2005). In pre-industrial development stage, as many countries experienced a low level of income per capita, the environment tended to be well preserved (Everett, Ishawaran, Ansaloni, & Rubin, 2010). The EKC dessibes the degradation of an environment as an inverted U-shape function to explain income per capita (Huang & Labys, 2002). In pre-industrial development stage, as many countries experienced a low level of income per capita, environment tended to be well preserved (Everett et al., 2010). However, the reverse happens when income per capita increases following the development of economic activities.

This situation could be explained using EKC hypothesis of Copeland and Taylor (2004) covering three important issues. First, the increase of negative impact on the environment in developing countries in 27 e with pollution haven hypothesis, while the pollution haven hypothesis is only one of the evidences that affect trade and the 5 vironment. Second, revenue increase has a positive impact on the environment through changes in the economic structure. Third, national income grows in tandem with environmental awareness which is driven by

stringent environmental regulations and replacement of environmentally unfriendly technologies with clean and friendly ones.

These three important issues can be explained using the inverted U-shape curve of Panayotou (1993) as described below. First, foreign investment flows lead to structural transformation from agriculture to manufacturing. Second, rising income leads to demand for better environmental quality. In developing countries, low level of income per capita poses challenges for the government to protect the environment as a result of rapid industrialisation. However, rising income levels can mpensate the environmental damage resulting from economic activities. At this stage, people are willing to sacrifice the consumption of goods for protecting the environment (And eoni & Levinson, 2007).

There has been no standard environmental presction indicators. Some scholars such as Holtz-Eakin and Selden (1995); Roberts and Grimes (1997), use carbonglioxide emission as an indicator while Grossman and Krueger (1991); Panayotou (1997) use sulphur dioxide. The EKC hypothesis could be described using a linear, quadratic or cubic function (Panayotou, 1994).

### Trade Openness and Environmental Degradation

The earliest work on trade and environment is the classical theory of trade which is based on the perfect competitive model (Baumol & Oates, 1993; Krutilla, 1991; Taylor & Copeland, 2004 ). Several studies apply Ricardian trade model or models that combine the emissions level into Heckscher – Ohlin (HO).

In economic theory, the relation between rade liberalisation and the environment, called the pollution haven effect, is the most controversial point of debate among Pholars, as they are yet to reach a consensus on the presence or absence of the pollution haven effect. Baumol and Oates (1993) identify the pollution haven effect as the trade liberalisation among countries having different environmental protection standards and practices; Taylor and Copeland (2004) opine a country should be responsible forpaying the social cost of environmental protection. Trade between countries usually does not involve environmental issues and many developing countries pay a high environmental cost to gain commodity comparative advantage. Therefore, attempts to override the effects of the polluting impact of trade as a possible spill over effect may exacerbate other distortions, namely low production and income levels due to low technological mastery in developing countries, and hence, need comprehensive policy reforms (Antweiler, Copeland, & Taylor, 2001).

### METHODS

This study employed time soies data from 1975 to 2014 of five Southeast Asian countries, including Indonesia, Malaysia, Singapore, Thailand, and the Philippines. The CO<sub>2</sub> emissions representing

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environmental degradation were used as the dependent variable. The economic condition represented by the real GDP per capita and the openness represented by trade ratio to GDP was the independent variables.

### The ECM model of EKC of five Southeast Asian countries

To examine the EKC hypothesis, this study computes the three of regression equation models as follows:

Linear : 
$$Y_t = \beta_0 + \beta_1 X_t + \varepsilon_{it}$$
 (1)

Quadratic : 
$$Y_t = \beta_2 + \beta_3 X_t + \beta_4 (X_t)^2 + \beta_5 T L_t + \varepsilon_{2t}$$
(2)

Cubic : 
$$Y_t = \beta_6 + \beta_7 X_t + \beta_8 (X_t)^2 + \beta_9 (X_t)^3 + \beta_{10} T L_t + \varepsilon_{3t}$$
 (3)

*Y<sub>t</sub>* is CO<sub>2</sub> emissions in year t (in kg/ton); *X<sub>it</sub>* is the reat<sub>25</sub>DP per capita in year *t* (in US\$); and *TL<sub>t</sub>* is the ratio of exports to GDP in year *t* (in percent) as the trade liberalisation indicator. The  $\beta_0$ ,  $\beta_2$ , and  $\beta_3$  are constants in equations (1), (2) and (3) successively;  $\beta_1$ ,  $\beta_3$ ,  $\beta_4$ ,  $\beta_5$ ,  $\beta_7$ ,  $\beta_8$ ,  $\beta_9$ , and  $\beta_{10}$ , are the regression coefficients and  $\varepsilon_{1b}$   $\varepsilon_{2b}$   $\varepsilon_{3t}$  are the error terms for equation (1), (2) and (3) successively in year *t*.

This study applied Engle-Granger (**†**G) procedure of Error Correction Model (ECM) to analyse the dynamics of the short-run effect of economic growth on environmental degradation, where the long-run relationships can be performed by estimating the regression of all cointegrated equations (**†**)omas, 1997). The co-integration of the two (or more) time series variables indicates there is a longrun nexus or equilibrium between the variables, then the short-run disequilibrium relationship between dependent and independent variables can be estimated.

EG test is conducted to test the cointegration of the residual  $\varepsilon_{3t}$ , using the Dicky-Fuller (DF) stationary test (Gujarati & Porter, 2009):

$$\Delta u_t = \Omega_1 \varepsilon_{3t-1} + v_t \tag{4}$$

The following equations are a brief derivation of ECM model based on equation (3); while, equation (1) and equation (2) are explained using similar procedures. To estimate the ECM model, equation (3) needs to be re-ordered by first adding the lag of independent variables, and the model is autoregressive distributed lag (ARDL) (1,1):

$$Y_{t} = b_{0} + b_{1}X_{t} + b_{2}X_{t-1} + b_{3}(X_{t})^{2} + b_{4}(X_{t-1})^{2} + b_{5}(X_{t})^{3} + b_{6}(X_{t-1})^{3} + b_{7}TL_{t} + b_{8}TL_{t-1} + \mu Y_{t-1} + e_{t}$$
(5)

and the result is,

$$\Delta Y_{t} = b_{1} \Delta X_{t} + b_{3} \Delta X_{t}^{2} + b_{5} \Delta X_{t}^{3} - (1 - \mu) (Y_{t-1} - \frac{b_{0}}{1 - \mu} - \frac{(b_{1} + b_{2})}{1 - \mu} X_{t-1} - \frac{(b_{3} + b_{4})}{1 - \mu} (X_{t-1})^{2} - \frac{(b_{5} + b_{6})}{1 - \mu} (X_{t-1})^{3}) + e_{t}$$
(6)

From the ADRL (1,1), it is noted that where  $\beta_6 = \frac{b_0}{1-\mu}$ ,  $\beta_7 = \frac{(b_1+b_2)}{1-\mu}$ ,  $\beta_8 = \frac{(b_8+b_4)}{1-\mu}$ , and  $\beta_9 = \frac{(b_5+b_6)}{1-\mu}$ . Equation (6) can be referred to a first order model ECM equation with  $\left(Y_{t-1} - \frac{b_0}{1-\mu} - \frac{(b_1+b_2)}{1-\mu}X_{t-1} - \frac{(b_4+b_4)}{1-\mu}(X_{t-1})^2 - \frac{(b_5+b_6)}{1-\mu}(X_{t-1})^3\right)$  as an *Error Correction Term*.

## Panel Data Regression Model of EKC in Southeast Asian Countries

Fixed Effects Model (FEM) and Random Effects Model (REM) are selected for the fit model used in this study. In this case, FEM assumes the intercept of each individual to have differences caused by special characteristics possessed by each individual. The terminology of the fixed effect shows that despite the fact the intercept varies among the individuals, for each individual, the intercept does not vary over time, called time invariant. It can also be stated that based on FEM models, it is assumed the slope coefficient of regressors does not vary between individuals and over time. The REM intercept is considered as the average of random variables. The intercept is not assumed to be constant, so the model is also popularly known as random error component model (Gujarati & Porter, 2009).

The panel regression equations of the hypothesis of EKC are formulated as follows (in the form of logarithm *Ln* quadratic models):

$$LnCO2_{it} = \beta_0 + \beta_1 LnGDPCAP_{it} + \beta_2 (LnGDPCAP_{it})^2 + \beta_3 LnTL_{it} + u_{it}$$
(9)

To estimate the EKC hypothesis, the  $CO_2$ emissions is used as a representation of environmental degradation, the Gross Domestic Product per Capita (*GDPCAP*) and the ratio of exports to gross domestic product (TL) are used to represent the macroeconomic condition and trade liberalisation. The subscript i indicate countries.

The fixed-effect model treats the  $\beta 0$  and  $\beta_t$  as regression parageters as they helped capture the countries' specific time invariant factors affecting pollution intensity. The random-effect model treats  $\beta_0$  and  $\beta_t$  as components of random disturbances. If the correlation between  $\beta_0$ ,  $\beta_t$  and the explanatory variables exist, the random-effect model cannot be estimated consistently (Hsiao,

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2007). Only the fixed-effect model can be estimated consistently. The Hausman test will be used to determine whether the fixed effect or random effect model is preferred.

### RESULTS AND DISCUSSION

For the long run models (Table 1),  $R^2$  values explain the variation of  $CO_2$  emissions more than (2) percent in each of estimation models of Indonesia, Malaysia, Thailand, and the Philippines, except Singapore (the  $R^2$  is 56.4). The figure of the calculated F was found greater than critical F, which pointed to the overall independent variables affecting  $CO_2$  emissions in each country (Table 1).

The long run estimation results indicate the models of EKC inverted U-shaped were empirically proven in Singapore, indicated by the coefficient  $\beta_1 > 0$  and  $\beta_2$ <0. Meanwhile, in Indonesia, the model generated  $\beta_1 > 0$  and  $\beta_2 > 0$  exhibiting an increase in emissions CO2 along with the increase of its economic growth; although, its per capita income does not statistically significant affect the CO2 at 10 percent level of confidence ( $\alpha$ ). Moreover, the estimation model of Malaysia, Thailand, and the Philippines are  $\beta_1 < 0$  and  $\beta_2 > 0$  meaning that CO<sub>2</sub> emissions decrease when per capita income increases, despite at a certain point, the per capita income increases CO<sub>2</sub> emissions.

### 1 Table 1

Estimation result of long run EKC model: Dependent variable Y

Variable	Coefficient and (t-stat)							
	Indonesia	Malaysia	Thailand	Philippines	Singapore			
Constant	3.102	34.843	9.050	51.009	-105.567			
	(0.552)	(2.978)***	(1.502)*	(1.521)	(-5.938)***			
LnX	0.707	-8.118	-1.387	-13.781	23.494			
	(0.413)	(-2.767)***	(-0.850)	(-1.453)	(6.650)***			
$LnX^2$	0.062	0.596	0.163	1.076	-0.120			
	(0.497)	(3.424)***	(1.435)*	(1.610)	(0.524)***			
LnTL	0.298	0.509	0.826	0.829	0.229			
	(2.193)**	(3.118)***	(1.502)***	(14.020)***	(0.524)			
$\mathbb{R}^2$	0.971	0.985	0.979	0.933	0.563			
Adj. R <sup>2</sup>	0.969	0.984	0.977	0.927	0.527			
F-stat	405.672	787.6747	549.756	165.891	15.501			

*Note:* \*\*\* significant at  $\alpha = 0.01$ ; \*\* significant at  $\alpha = 0.05$ ; \* significant at  $\alpha = 0.10$ 

The estimations of the long 3n model showed trade liberalisation had a positive and significant effect on CO<sub>2</sub> emissions, except in Singapore. Statistically, trade liberalisation in the island state did not significantly affect CO<sub>2</sub> emissions but it pointed to a positive relationship. This finding proves that Singapore economy as a developed country, places great concern on the quality of its environment.

According to ECM estimation (Table 2), the short run model is consistent with the long run one. For Indonesia, the variable of trade liberalisation do not statistically significant affect environmental degradation. For all countries, except the Philippines, the

quadratic per capita income significantly affected CO<sub>2</sub> emissions. Based on the estimated coefficients of each country, the EKC hypothesis had only been occurred in Singapore, where  $\beta_1 > 0$  and  $\beta_2 < 0$ . In the short run, the trade liberalisation variable had significant effect ( $\alpha = 0.01$  and  $\alpha = 0.05$ ) on the Philippines and Malaysia estimation models while the Thai model showed a significant effect at  $\alpha = 0.1$ . The negative sign of FGT explains the discrepancy between the long-run and the short-run estimated CO<sub>2</sub> emissions was corrected by the magnitude of the ECT coefficient within one period.

1 Table 2

Estimatio	on result	of si	hort i	run	EKC	model:	Depena	lent	variable	$\Delta Ln Y$

Variable			Coefficient and (t-	stat)	
	Indonesia	Malaysia	Thailand	Philippines	Singapore
Constant	0.046	0.067	0.070	0.023	0.029
	(2.35)**	(4.279)***	(5.164)***	(2.202)	(0.596)
$\Delta LnX$	1.385	-8.816	-0.873	-9.696	14.553
	(0.291)	(-1.838)*	(-0.312)	(-0.864)	(1.356)
$\Delta LnX^2$	-0.067	0.522	0.045	0.709	-0.789
	(-0.193)	(1.794)*	(0.240)	(0.885)	(-1.457)
$\Delta LnTL$	-0.006	0.635	0.209	0.194	0.743
	(-0.062)	(3.194)**	(1.605)*	(1.503)***	(-3.132)
ECT	-0.370	-0.0535	-0.284	-0.422	-0.406
	(-3.224)***	(5.058)***	(-3.806)***	(-4.450)***	(-3.132)***
$\mathbb{R}^2$	0.256	0.481	0.305	0.387	0.268
Adj. R <sup>2</sup>	0.169	0.420	0.223	0.315	0.181
F-stat	2.930	7.874	3.732	5.370	3.106

*Note:* \*\*\* significant at  $\alpha = 0.01$ ; \*\* significant at  $\alpha = 0.05$ ; \* significant at  $\alpha = 0.10$ 

To find out how environmental issues affect Southeast Asian economies, this study applies panel data regression model, which is FEM - which is resulted by the Hausman model selection. The24 timation results indicate that coefficient of GDP per capita (in linear and quadratic forms) affect  $CO_2$  emissions in southeast Asian countries. The coefficient of per capita income and quadratic per capita income are statistically significant affecting CO<sub>2</sub> emissions at level of confidence  $\alpha = 0.01$  (Table 3). The coefficient 10 capita income of 4.830 points to the fact in the early stages of Southeast ian economic development, 1 percent increase in per capita income increases CO2 emissions by 4.830 percent.

The estimation of FEM model proves that a priori positive sign of the coefficient per capita income ( $\beta_1$ ) and the negative sign of quadratic per capita income ( $\beta_2$ ), form the inverted U curve of EKC hypothesis (indicated by  $\beta_2 \neq 0$ , and if  $\beta_1 > 0$  and  $\beta_2$ < 0). This finding means that panel data estimation was able to explain the behaviour of CO<sub>2</sub> emissions in the five Southeast Asian countries.

#### Table 3

Estimation results of the data panel - FEM

Variable	Coefficient	t-Statistic
Constant	-14.907	-12.396***
LnGDPCAP	4.830	15.419***
$LnGDPCAP^{2}$	-0.236	-12.758***
$T_{12}TL$	0.638	7.868***
R-squared	0.922	
Adjusted R-squared	0.920	
F-statistic	326.172	

Note: \*\*\* significant at  $\alpha = 0.01$ 

### CONCLUSION

This study utilised the dynamic econometric model and the cointegration test to estimate the long-run relationship of trade liberalisation on environmental degradation to prove the EKC hypothesis in the five Southeast Asian countries. In the long run, the estimations indicated that a priori model of EKC was supported in Singapore, and the trade liberalisation had positive significant effect on CO2 emissions, except in Singapore. In the short run, the estimation results proved the estimated parameters supported the EKC hypothesis, though the effect of income per capita and trade liberalisation on environmental degradation was not statistically significant.

The result shows the differences of national income per capita and trade liberalisation in relation to environmental degradation in eveloped country i.e. Singapore and in developing countries such as Indonesia, Malaysia, Thailand, and the Philippines in Southeast Asia. This could be due to the dependence of the latter countries on natural resources, the increase in the economics of scale, and the shift of the economic structure from agriculture to manufacture in four Southeast Asian countries, except Singapore.

The estimation of panel data regression of Southeast Asian countries indicated per 21 ita income and trade liberalisation had significant effect on CO<sub>2</sub> emissions. Therefore, it can be concluded that the environmental policies were crucial and must be coordinated in order to reduce their negative impacts on the environment due to

rapid development and expansion of trade. Meanwhile, the EKC hypothesis was proven in Southeast Asian countries. The increase in per capita income aggravate environmental degradation in Southeast Asian countries, but the turning point is the increase in income that would reduce environmental degradation due to awareness and the cost of the environmental damages.

Although the findings, to some extent, are conclusive, this study has several limitations including lack of data in a long-time series and the use of environmental indicators, which were limited to  $CO_2$  emissions. This study can be developed using other environmental indicators.

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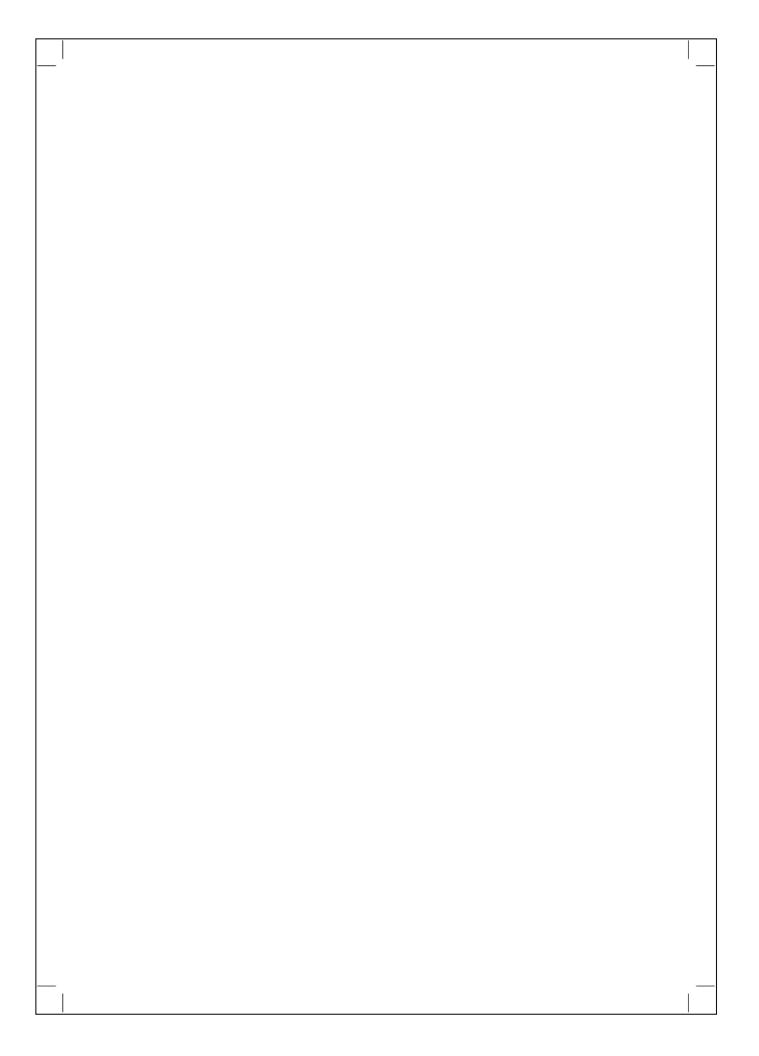
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